

1                    1                    **Measuring sexual selection in mammals**

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**Abstract**

Sexual selection has been increasingly acknowledged as a major evolutionary driver in many taxa. An unequal mating success arising through competition over mates is the phenomenon that defines sexual selection. Therefore, demonstrating the presence of sexual selection requires that the number or quality of mates are spread less evenly than random processes would suggest. The objective of this study was to estimate the extent of this inequality in male mammals by reviewing studies that measure variance in genetic paternity. I found that few species showed high values of the standardised variances in male reproductive success,  $I_m$ , and Nonacs's  $B$  indices. These species differed from those with low or medium values in that, while in the latter the females breed solitary, in the former, they are grouped for reproduction, which possibly facilitates mate monopolization by males. Low levels of variance in reproductive success could be explained because most species of mammals breed solitary or due to high rates of extra-pair copulations by 'subordinate' males.

**Keywords:** Evolution, paternity, genetic reproductive success, breeding system, mating system.

## 28 Introduction

29 Sexual selection is considered one of the most pervasive forces that drive the evolution  
30 of reproductive traits in sexually reproducing organisms (Andersson, 1994). Darwin  
31 (1871) defined sexual selection as being dependent “on the advantage which certain  
32 individuals have over other individuals of the same sex and species, in exclusive  
33 relation to reproduction” (p. 871). A modern definition proposes that sexual selection  
34 favours investment in traits that improve the likelihood of fertilization given limited  
35 access to opposite-sex gametes due to competition with members of the same sex  
36 (Jennions and Kokko, 2010). Sexual selection should be particularly intense in  
37 mammals because females bear most parental investment due to gestation and lactation,  
38 transforming them in a limiting resource for which males compete (Trivers 1972).

39 An unequal mating success that arises through competition over mates is the  
40 phenomenon that defines sexual selection (Arnold and Wade, 1984). Therefore,  
41 demonstrating the presence of sexual selection requires that the number or quality of  
42 mates are spread less evenly than random processes would suggest, in at least one of the  
43 sexes (Kokko et al., 2008). Kokko et al., (1999, p. 358) stated that, "unless the  
44 inequality itself can be shown to exist, and preferably measured and tested for  
45 significance, any further study of the causes behind it is fruitless".

46 In order to measure variance in mating success, a type of sexual selection index  
47 must be selected and the data that this type of index requires must be obtained. In the  
48 literature, there is a several indices (Klug et al., 2010, Anthes et al., 2017). Nevertheless,  
49 most indices correlate very strongly and measuring the same data with multiple indices  
50 will rarely reveal novel insights (Nonacs 2003). In this paper, I used the variance in  
51 relative reproductive success, the so called intensity (or opportunity) of sexual selection  
52 on males  $I_m$ , which is calculated as the variance divided by the square of the mean

(Arnold and Wade, 1984). I used  $I_m$  because it is the most common index used in mammalian literature (Isvaran and Sankaran 2017). Based on the range shown for mammalian species, Dubuc et al. (2014) considered values of  $I_m < 1$  to show weak opportunity for selection, values between 1 and 3 show moderate opportunity and values of  $I_m > 3$  show high opportunity for selection.

Ideally, data to measure sexual selection should be generated with an experimental design that includes the entire reproductive period (Fritzche and Booksmythe, 2013), because experiments can assess causal relationships (Anthes et al., 2017). Clearly, this type of experiment is impossible to conduct with most mammals. Thus, most of the classical studies in this taxon were conducted using behavioural traits that are indirectly related to reproductive success, e.g. harem size (Emlen and Oring 1997). Using these behavioural surrogates, the overall result was that mammals showed high skew in reproductive success (review, for example, in Clutton-Brock, 1988). Recently, these measures have been replaced by estimations of paternity using genetic techniques, which provides a valuable measure of the opportunity of sexual selection because it excludes the possibility that males who are not dominant in direct competition for mates may obtain extra-group copulations (Wade and Shuster, 2004).

The objectives of this work were: (i) to review the literature on genetic paternity in mammals in order (ii) to obtain estimates of the opportunity of sexual selection in this taxon, (iii) to compare genetic with behavioural measures of male reproductive success and (iv) to discuss the role of sexual selection in mammalian evolution.

## Methods

### *Creating datasets*

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77 For paternity data, the methodology of Isvaran and Sankaran (2017) was adopted which  
78 aimed to isolate estimates of access to fertilization (i.e., the effect of sexual selection on  
79 variation in reproductive success) and remove the effect of factors related to viability  
80 selection (i.e., survival between breeding seasons and lifespan). A literature search was  
81 conducted for estimates of male genetic mating success, based on genetic measures of  
82 paternity (Supplementary Information 1). This search had four stages. In the first, I used  
83 two keywords in Google Scholar database, 'paternity' and 'mammal', to get a first list of  
84 articles. The second stage consisted of a review of those journals that appeared in the  
85 first stage, using the same two words in the search tools available in the web pages of  
86 the journals themselves. The third stage consisted of a review of the literature cited in  
87 all the articles obtained in the second stage. The fourth stage consisted of a review of  
88 the literature cited in all the articles obtained in the third stage. Only fieldwork studies  
89 were considered unless otherwise mentioned.

90 In order to obtain a measure of inequality in reproductive success, it was  
91 necessary to extract the information on paternity from each study. In all studies,  
92 paternity was established through the use of molecular markers and, in order to do so,  
93 skin or blood samples were taken from those parents and offspring that were captured  
94 and then the degree of kinship was determined. In most cases, male reproductive  
95 success was presented as the number of pups that were assigned to each sampled male.  
96 Therefore, I calculated the variance and the mean in male reproductive success of the  
97 total of the sampled males. In some cases, the variance and the mean were directly  
98 provided by the study. In studies that report male reproductive measures counted across  
99 multiple years, the annual mean number of offspring per male was obtained first, to then  
100 obtain the mean and variance for the total number of males.

Female breeding system data were obtained from Lukas and Clutton-Brock (2020). Information gaps were filled with information provided by the Animal Diversity Web, University of Michigan (<https://animaldiversity.org/> accessed September - November 2020). Data on mating systems were obtained from the same sources of paternity data.

### ***Measuring inequality in reproductive success***

I used the 'opportunity of sexual selection' calculated as  $I_m = var/mean^2$  of paternity data. I followed Dubuc et al. (2014)'s criteria to define low, medium and high skew in male reproductive success. Complementing and for comparison, I calculated the Nonacs' *B* index (Nonacs, 2000, 2003) using the program Skew Calculator 2003 (<http://www.eeb.ucla.edu/Faculty/Nonacs/PI.html>), which tests whether paternity distribution is different from that expected by chance. This program allows splitting into separate groups, but I did not so because only a few number of studies provided data by group. This program provides the following confidence intervals of *B*: (i) if they include zero, the distribution of benefits is not different from random chance; (ii) if the lower confidence interval includes the 'equal sharing' value (also provided), then an equal distribution of benefits cannot be excluded; (iii) if the intervals do not include zero, this indicates a skew distribution of paternity, with some males being significantly more successful than others; (iv) under the conditions of (iii), *B* values close to zero indicate low variance in mating success, while *B* values close to one indicates high levels of paternity monopolisation. Therefore, values  $< 0.5$  were considered low skew values, and  $> 0.5$  were considered high. I included studies that both sampled and not sampled unsuccessful males. This could affect the level of skew estimated with  $I_m$  but is more controlled by *B*.

**Relating variables**

All relationships were analysed employing the phylogenetic generalised least squares (PGLS) method in R, so the phylogenetic signal could be controlled (Martins and Hansen 1997). First, a simple regression was conducted to compare  $I_m$  with  $B$ , then a multiple regression was conducted with  $I_m$  as a dependent variable (log transformed, with five zero-data assigned a value of -2) and four independent variables: body size (log transformed), sexual size dimorphism (log transformed) and two proxy measures of strength of sexual selection, based on the intensity of the male competition for mates: mating system (polygyny, polygynandry, monogamy) and female breeding system (plural versus single breeders).

**Results**

A dataset was built containing genetic variance information from a total of 158 mammal species, of which 156 had data that allowed the estimation of  $I_m$ , and 126 had data that allowed the estimation of  $B$ . Mean  $I_m$  and  $B$  were 1.48 (SD=2.04) and 0.027 (SD=0.098). As expected,  $I_m$  and  $B$  were strongly correlated ( $\lambda=0$ ,  $r^2=0.22$ ,  $F_{1, 116}=33.8$ ,  $p<0.001$ ). According to the categorisation made by Dubuc et al. (2014), most species showed weak selection opportunity (50.3%), while only 13.8% of species had high levels of selection opportunity and 35.8% showed a moderate opportunity for sexual selection (Figure 1). These frequencies showed statistically significant differences ( $\chi^2=32.2$ ,  $gl=21$ ,  $P<0.001$ ). The  $B$  index showed a similar pattern (Figure 1), given that the number of species with non-skewed (equal or random) distributions (59.1%) was higher than the number of species with a skewness ( $\chi^2=4.2$ ,  $gl=1$ ,  $P=0.04$ ).

Four behavioural measures of sexual selection were compared with  $I_m$  (Table 1).

The only significant correlation was between  $I_m$  and the breeding system.

## Discussion

Results obtained in this study, which used the largest available dataset on genetic measures of paternity variance in mammals, suggest low levels of opportunity of sexual selection in this taxon. A possible explanation of these low values of genetic measures of variance in reproductive success would be the effect of extra-group copulations. Most mammalian species were described as behaviourally polygynous, with consequent variance in the ability of males to monopolize females (Emlen and Oring, 1977). Large males can monopolise females by contest competition, while small males can evolve alternative reproductive tactics, such as sneak spawning (Jennions and Kokko 2010). If extra-pair copulations are achieved by less successful males in direct competition, the result is a decrease in the variance in male reproductive success. In populations with more than one type of male reproductive tactic, it could be that ethological observations indicate a high variance in reproductive success with large males gaining more mates than small ones, but this does not necessarily imply a high paternity variance, since alternative tactics could be equally effective but less conspicuous, with the consequent homogenization of this variable. Under this hypothesis, sexual selection operates maintaining equilibrium between two or more alternative reproductive strategies depending on their size or age.

This hypothesis was partially tested by Isvaran and Sankaran (2017), who showed that when breeding group sex ratios are strongly female-biased (highly polygynous systems, multi-male systems with female-biased groups), extra-group



183 175 paternities appear to decrease variance in male reproductive success. These authors  
184 176 described this result as 'an unsuspected relationship between overt (social mating  
185 177 system) and covert (extra-group paternity) mating tactics and the potential for sexual  
186 178 selection in male mammals'.

187 179 A correlation between breeding system and intensity of sexual selection was  
188 180 found. This relationship is predicted by sexual selection theory, which postulates that  
189 181 the extent to which breeding females aggregate affects mating competition (Emlen and  
190 182 Oring, 1977). In plural breeding species, the number of potential breeding partners that  
191 183 individual males can guard effectively is higher than in singular breeders, increasing  
192 184 variance in male reproductive success (Kvarnemo and Ahnesjo, 1996). In most  
193 185 mammalian species, females breed solitary (Lukas and Clutton-Brock 2020), therefore it  
194 186 is expected relatively low levels of intensity of sexual selection in this taxon if sexual  
195 187 selection increases with female group size.

196 188 Another hypothesis is that current variation in reproductive success refers only  
197 189 to the maintenance of sexually selected traits, not their origin and further evolution in  
198 190 the past, when variance in reproductive success might have been different (Grafen  
199 191 1988). In other words, the trait would cause skewed reproduction among males, but  
200 192 would quickly become fixed (Fisher's run-away model of sexual selection, Fisher 1930).

201 193  
202 194 This study only analyses the current opportunity of sexual selection measured as  
203 195 variance in parenting but does not say anything about the impact of sexual selection on  
204 196 actual traits, such as sexual size dimorphism. Those traits need to be heritable and have  
205 197 strong enough fitness advantages to show change in trait forms over evolutionary time.  
206 198 Therefore, a future study should analyse the relationship between sexual traits and  
207 199 paternity variance in mammals.

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263 252 male reproductive success. *The American Naturalist* 164(4): E83-E89.
- 264 253

266 254 Table 1. Results of the phylogenetic generalised least squares using standardised  
 267 255 variance in male reproductive success  $I_m$  as dependent variable and four independent  
 268 256 variables: breeding and mating systems. The global analysis gave:  $\lambda=0$ ,  $r^2=0.08$ ,  $F_{1,}$   
 269 257  $_{133}=2.94$ ,  $p=0.02$   
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Independent variable	Estimate	Std. Error	t-value	p-value
(Intercept)	-0.45	0.15	-30.8	0.0002
Body mass	0.01	0.04	0.14	0.89
Sexual dimorphism	0.43	0.40	1.06	0.29
Breeding system	0.33	0.11	29.4	0.003
Mating system	0.02	0.04	0.64	0.52

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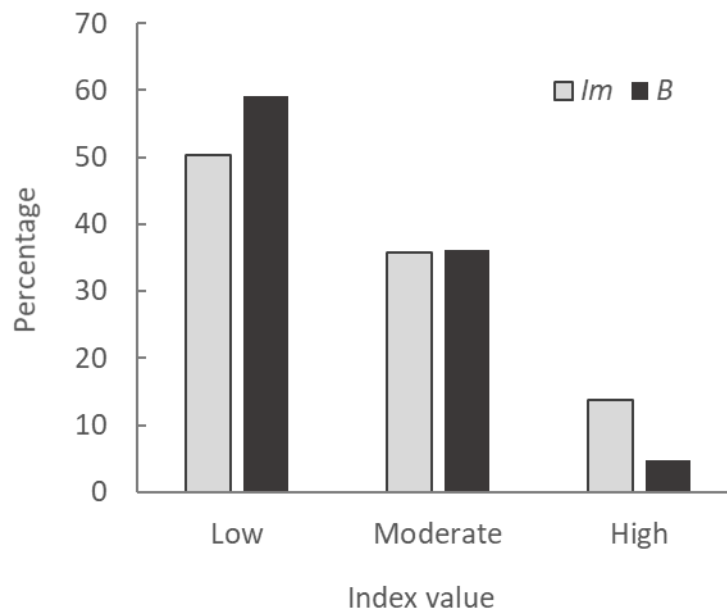
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261 Figure 1. Percentages of low, moderate and high values of standardised variance in male  
262 reproductive success was calculated:  $I_m$  and Nonacs'  $B$  index. Low values of  
263 opportunity of sexual selection were significantly more frequent than high values.  
264  $\chi^2=32.2$ ,  $gl=21$ ,  $P<0.001$



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